# On the strong convergence of the $\varepsilon$ -EM scheme for time-inhomogeneous jump driven SDEs

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<sup>1</sup> This is a joint work with Mireille Bossy.





#### Introduction : additive jump models and applications

▶ We are interested in the approximation and the simulation of the solution process  $(X_t, t \in [0, T])$  of the SDE

$$X_t = \int_0^t b(X_s) ds + \int_0^t \int_{-\infty}^\infty c(s, X_{s^-}, z) \widetilde{N}(ds, dz),$$

where  $\widetilde{N}$  is a compensated **random Poisson measure**, with **time-inhomogeneous** compensator measure  $v_s(dz)ds$ , and b and c are deterministic Lipschitz-in-space coefficients.

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- This class of SDE is useful to obtain non-gaussian stochastic models that may have several time regimes. Such a model can be for example used
  - to describe the population dynamics of parasitoid insects (see [BCP<sup>+</sup>23])
  - to capture option prices over a range of different maturities and strikes (see [CT04])
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  - to capture option prices over a range of different maturities and strikes (see [CT04])
  - to represent the angle dynamics of non-spherical particles in a turbulent flow (see [CBB22]).
- Our goal is to construct an algorithm (X̄<sub>n</sub>) to simulate the process (X<sub>t</sub>) and obtain rates of convergence for the probabilistic strong error (trajectorial) between X̄ and X.

Numerical algorithms to approximate SDEs driven by a Brownian motion W are well known in the literature:

$$X_t = \int_0^t b(X_s) ds + \int_0^t \sigma(X_s) dW_s.$$

The simplest method is given by the Euler-Maruyama scheme:

$$\overline{X}_{t_{i+1}} = \overline{X}_{t_i} + b(\overline{X}_{t_i})(t_{i+1} - t_i) + \sigma(\overline{X}_{t_i})(W_{t_{i+1}} - W_{t_i}),$$

where  $(t_i = \frac{iT}{n}, i = 0, ..., n)$  are the discretization steps, and  $n \in \mathbb{N}$ .

This scheme is straightforward to implement because one knows how to simulate the law of any increment W<sub>s</sub> - W<sub>u</sub> of the Brownian motion easily.

A first generalisation consists in replacing the driving Brownian motion by a Lévy process L, i.e a process L having independent and stationary increments:

$$X_t = \int_0^t b(X_s) ds + \int_0^t \sigma(X_{s^-}) dL_s.$$

In this case, one may still define the Euler-Maruyama scheme:

$$\overline{X}_{t_{i+1}} = \overline{X}_{t_i} + b(\overline{X}_{t_i})(t_{i+1} - t_i) + \sigma(\overline{X}_{t_i})(L_{t_{i+1}} - L_{t_i}),$$

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However, simulating the increments L<sub>s</sub> – L<sub>u</sub> of the Lévy process is not something easy in general. It is possible for some particular example, such as the so called α-stable process L<sub>α</sub>.

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- However, simulating the increments L<sub>s</sub> L<sub>u</sub> of the Lévy process is not something easy in general. It is possible for some particular example, such as the so called α-stable process L<sub>α</sub>.
- Thanks to the Lévy-Itô décomposition, if L has at least a moment of order 2, it can be written

$$L_t = \int_0^t \int_{-\infty}^\infty z(N(ds, dz) - \nu(dz)ds) = \int_0^t \int_{-\infty}^\infty z\widetilde{N}(ds, dz) ds$$

where v is a deterministic measure called the Lévy measure of L.

A second generalisation consists in replacing the driving Brownian motion by an additive process A, i.e a process A having independent increments:

$$X_t = \int_0^t b(X_s) ds + \int_0^t \sigma(X_{s^-}) dA_s.$$

In this case, one may still define the Euler-Maruyama scheme:

$$\overline{X}_{t_{i+1}} = \overline{X}_{t_i} + b(\overline{X}_{t_i})(t_{i+1} - t_i) + \sigma(\overline{X}_{t_i})(A_{t_{i+1}} - A_{t_i}),$$

where  $(t_i = \frac{iT}{n}, i = 0, ..., n)$  are the discretization steps, and  $n \in \mathbb{N}$ .

- However, simulating the increments A<sub>s</sub> A<sub>u</sub> of the additive process is not something easy in general.
- Thanks to the Lévy-Itô décomposition, if A has at least a moment of order 2, it can be written

$$A_t = \int_0^t \int_{-\infty}^\infty z(N(ds, dz) - \mathbf{v}_s(dz)ds) = \int_0^t \int_{-\infty}^\infty z\widetilde{N}(ds, dz)$$

where  $(v_s)_{s \in [0,T]}$  is a collection of Lévy measures.

Finally, based on the Lévy-Itô decomposition, one may think of an "increments-free" generalisation, leading to the SDE:

$$X_t = \int_0^t b(X_s) ds + \int_0^t \int_{-\infty}^\infty c(X_{s^-}, z) \widetilde{N}(ds, dz).$$

However, in this case we are not able to give a sense to the Euler-Maruyama scheme:

$$\overline{X}_{t_{i+1}} = \overline{X}_{t_i} + b(\overline{X}_{t_i})(t_{i+1} - t_i) + ???$$

where  $(t_i = \frac{iT}{n}, i = 0, ..., n)$  are the discretization steps, and  $n \in \mathbb{N}$ .

Indeed, we can't define a discretisation of X that relies on the increments of an underlying stochastic process anymore, except in the case where c can be written c(x, z) = σ(x)f(z).

► Let  $F: [0,T] \times \mathbb{R} \to \mathbb{R}$  be a measurable function such that  $\int_0^T \int_{-\infty}^{\infty} |F(s,z)|^2 v_s(dz) ds < \infty$ . We want to simulate the stochastic integral:

$$I(F) = \int_0^T \int_{-\infty}^{+\infty} F(s,z) \widetilde{N}(ds,dz)$$

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$$I(F) = \int_0^T \int_{-\infty}^{+\infty} F(s, z) \widetilde{N}(ds, dz)$$

► Taking the threshold ε > 0, I(F) can be separated into it's large jumps part I<sup>ε</sup><sub>l</sub>(F) and small jumps part I<sup>ε</sup><sub>l</sub>(F):

$$I_{l}^{\varepsilon}(F) = \int_{0}^{T} \int_{\mathbb{R}\setminus B(\varepsilon)} F(s,z)\widetilde{N}(ds,dz), \qquad I_{l}^{\varepsilon}(F) = \int_{0}^{T} \int_{B(\varepsilon)} F(s,z)\widetilde{N}(ds,dz),$$
$$I(F) = I_{l}^{\varepsilon}(F) + I_{l}^{\varepsilon}(F).$$

#### The large jumps : a direct simulation method

The large jump integral can be represented by the difference of a finite random sum and a deterministic integral:

$$\int_{0}^{T} \int_{\mathbb{R}\setminus B(\varepsilon)} F(s,z) N(ds,dz) = \sum_{j=1}^{N^{\varepsilon}(T)} F(T^{\varepsilon}(j), Z^{\varepsilon}(j)) - \int_{0}^{t} \int_{\mathbb{R}\setminus B(\varepsilon)} F(s,z) v_{s}(dz) ds.$$
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► In the above formula,  $N^{\varepsilon}$  is a (time-inhomogeneous) Poisson process with intensity function  $\lambda^{\varepsilon}(t) = \int_{\mathbb{R}\setminus B(\varepsilon)} v_t(dz)$ , and jump times  $T^{\varepsilon}(j) = \inf\{t \in [0,T], N^{\varepsilon}(t) = j\}$ .

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The random variables Z<sup>ε</sup>(j) for j ≤ N<sup>ε</sup>(T) have conditional distribution given the jump times given by:

$$\forall B \in \mathscr{B}(\mathbb{R}) \qquad \mathbb{P}(Z^{\varepsilon}(j) \in B \mid T^{\varepsilon}(j) = t) = \frac{\nu_t(B \cap \mathbb{R} \setminus B(\varepsilon))}{\nu_t(\mathbb{R} \setminus B(\varepsilon))} \ .$$

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▶ Note that if the time-dependence of  $v_t$  is multiplicative, i.e if one has  $v_t(dz) = \phi(t)v(dz)$ , then the latter distribution is homogeneous in time. In this case the jump sizes  $Z^{\varepsilon}(j)$  are i.i.d.

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- Hence, to perform a simulation of  $I_l^{\varepsilon}(F)$ , one needs:
- 1. To generate the Poisson process  $N^{\varepsilon}$ : this can be done with thinning method if the function  $\lambda$  is bounded, or by (eventually numerical) inversion of  $t \mapsto \lambda(t)$ ;
- To generate the jump sizes random variables Z<sup>ε</sup>(j) for any j ≤ N<sup>ε</sup>(T): for usual distributions, this can be done by inversion or acceptance-rejection methods;
- 3. To compute the deterministic integral  $\int_0^t \int_{\mathbb{R}\setminus B(\varepsilon)} F(s,z) v_s(dz) ds$ : it can be done analytically or numerically, depending on the difficulty.

#### The large jumps : a simple (and useful !) example

• We take the example of a 1-truncated  $\alpha$ -stable process, i.e F(s,z) = z and

$$\mathbf{v}_t(dz) = f(t)|z|^{-(1+\alpha)} \mathbb{1}_{\{|z| \le 1\}},$$

where  $\alpha \in (0,2]$  and  $f \in L^{\infty}([0,T])$ .

- 1. The Poisson process  $N^{\varepsilon}$  has intensity function  $\lambda^{\varepsilon}(t) \leq 2 \|f\|_{\infty} \frac{\varepsilon^{-\alpha} 1}{\alpha}$ , allowing to use a thinning method.
- 2. The jump sizes  $Z^{\varepsilon}(j)$  are i.i.d and  $Z^{\varepsilon}(1)$  has explicit quantile function given by

$$\forall y \in ]0,1], \quad \mathcal{Q}_{\mathcal{E}}(y) = \begin{cases} -\{2y(\mathcal{E}^{-\alpha}-1)+1\}^{-\frac{1}{\alpha}}, & \text{if } y \in (0,\frac{1}{2}], \\ \{(1-2y)(\mathcal{E}^{-\alpha}-1)+\mathcal{E}^{-\alpha}\}^{-\frac{1}{\alpha}}, & \text{if } y \in (\frac{1}{2},1]. \end{cases}$$

3. The deterministic integral  $\int_0^t \int_{\mathbb{R}\setminus B(\varepsilon)} F(s,z) v_s(dz) ds$  is straightforward to compute:

$$\int_0^t \int_{\mathbb{R}\setminus B(\varepsilon)} z \nu_s(dz) ds = \left( \int_0^t f(s) ds \right) \left( \int_{\varepsilon \le |z| \le 1} z |z|^{-1-\alpha} dz \right) = 0.$$

#### The small jumps : an extension of the Asmussen-Rosinski method

Generally, exact simulation of the small jumps integral I<sup>e</sup><sub>I</sub>(F) is not possible, but we may approximate it using the following idea:



<sup>1</sup> Obtained by direct simulation, which is possible in this very specific case thanks to the acceptance-rejection algorithm developped by Dassios, Lim and Qu in [DLQ19]

#### The small jumps : an extension of the Asmussen-Rosinski method

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- We substitute the stochastic integral I<sup>e</sup><sub>l</sub>(F) with a Gaussian random variable having an equivalent variance:

$$\mathscr{L}aw(I_l^{\varepsilon}(F)) \simeq \left(\int_0^T \int_{B(\varepsilon)} |F(s,z)|^2 v_s(dz) ds\right)^{\frac{1}{2}} \mathcal{N}(0,1),$$

where  $\mathcal{N}(0,1)$  designates the standard normal distribution.

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To quantify the error made in this approximation, we will use the Wasserstein distance of order p defined by

$$\mathscr{W}_p(\mathscr{L}_1, \mathscr{L}_2) = \inf_{(X_1, X_2) \in \pi(\mathscr{L}_1, \mathscr{L}_2)} \mathbb{E}[|X_1 - X_2|^p]^{\frac{1}{p}},$$

where  $(X_1, X_2) \in \pi(\mathscr{L}_1, \mathscr{L}_2)$ , means that the random variables  $X_1, X_2$  verify  $\mathscr{L}aw(X_i) = \mathscr{L}_i$ .

#### The small jumps : an extension of the Asmussen-Rosinski method

#### Proposition 1

Let  $p \ge 1$ . Assume that  $F(t, \cdot)$  is non identically zero on B(1) for any  $t \in [0, T]$  and

$$\int_0^T \int_{B(1)} |F(s,z)|^{p+2} v_s(dz) ds + \int_0^T \left( \int_{B(1)} |F(s,z)|^2 v_s(dz) \right)^{\frac{p}{2}+1} ds < \infty.$$

Then there exists a constant  $\mathscr{A}(p)$ , only depending on p, such that for every  $\varepsilon \in (0,1)$ , the following inequality holds for any  $t \in [0,T]$ :

$$\begin{aligned} & \mathscr{W}_{p}\left(\mathscr{L} \operatorname{aw}\left(\int_{0}^{t}\int_{B(\varepsilon)}F(s,z)\widetilde{N}(ds,dz)\right), \ \mathscr{N}\left(0,\int_{0}^{t}\int_{B(\varepsilon)}|F(s,z)|^{2}\nu_{s}(dz)ds\right)\right) \\ & \leq \mathscr{A}(p)\left(\frac{\int_{0}^{t}\int_{B(\varepsilon)}|F(s,z)|^{p+2}\nu_{s}(dz)ds}{\int_{0}^{t}\int_{B(\varepsilon)}|F(s,z)|^{2}\nu_{s}(dz)ds}\right)^{\frac{1}{p}}. \end{aligned}$$

$$(1)$$

The term in the right hand-side goes to zero when ε goes to zero on good conditions on F (a sufficient condition is that lim<sub>|z|→0</sub> sup<sub>s∈[0,T|</sub> |F(s,z)| = 0).

#### Idea of the proof

The proof relies on a W<sub>p</sub>-distance quantification of the convergence of the CLT (Rio's conjecture, proved by Bobkov in 2018 in [Bob18]):

#### Theorem – (S.G Bobkov, 2018)

For  $p \ge 1$ , there exists  $c_p > 0$  depending only on p such that if  $X_1, \ldots, X_m$  are independent random variables with  $\sum_{i=1}^m \operatorname{Var}(X_j) = 1$ , then

$$\mathcal{W}_p\left(\mathscr{L}\textit{aw}\left(\sum_{j=1}^m X_j\right), \ \mathcal{N}(0,1)\right) \leq c_p\left(\sum_{j=1}^m \mathbb{E}[|X_j|^{p+2}]\right)^{\frac{1}{p}}$$

We apply this result to the independent random variables

$$X_j = \int_{\tau_{j-1}}^{\tau_j} \int_{B(\varepsilon)} F(s,z) \widetilde{N}(ds,dz), \quad j \in \{1,\ldots,m\},$$

where  $\tau_j = \frac{jt}{m}$ , and estimate the p + 2-moment of  $X_j$  using Kunita inequality for random Poisson integrals, that we will recall later in this talk.

#### 3. The $\varepsilon$ -EM scheme

- This discussion allows us to define what we call the ε-Euler Maruyama scheme to approximate the process X<sub>t</sub> in introduction.
- We fix a threshold ε ∈ (0,1). For n ∈ N\*, we define 0 = t<sub>0</sub> < ··· < t<sub>n</sub> = T, a discretisation of the interval [0, T] with constant steps, i.e t<sub>i</sub> = i T/n. Let (ξ<sub>i</sub>)<sub>i∈{1,...,n}</sub> a sequence of i.i.d standard Gaussian random variables. We define X<sup>ε</sup> by X<sup>ε</sup><sub>0</sub> = 0 and

$$\begin{split} \overline{X}_{t_{i}}^{\varepsilon} &= \overline{X}_{t_{i-1}}^{\varepsilon} + b(t_{i-1}, \overline{X}_{t_{i-1}}^{\varepsilon}) \frac{T}{n} - \int_{t_{i-1}}^{t_{i}} \int_{\mathbb{R}\setminus B(\varepsilon)} c(s, \overline{X}_{t_{i-1}}^{\varepsilon}, z) v_{s}(dz) ds \\ &+ \left( \int_{t_{i-1}}^{t_{i}} \int_{B(\varepsilon)} c^{2}(s, \overline{X}_{t_{i-1}}^{\varepsilon}, z) v_{s}(dz) ds \right)^{\frac{1}{2}} \xi_{i} + \sum_{j=N^{\varepsilon}(t_{i-1})+1}^{N^{\varepsilon}(t_{i})} c(T^{\varepsilon}(j), \overline{X}_{t_{i-1}}^{\varepsilon}, Z^{\varepsilon}(j)). \end{split}$$

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► We will know be giving a convergence result for  $\overline{X}^{\varepsilon}$  in the  $L^p$ -norm. Note that this convergence will depend on two parameters, which are the number *n* of discretisation steps and the small jumps/big jumps threshold  $\varepsilon$ .

### 4. Hypothesis required for $L^p$ -strong convergence

We fix  $p \ge 2$  and T > 0. We consider a filtered probability space  $(\Omega, \mathcal{F}, \mathbb{P}, (\mathcal{F}_t)_{t \in \mathbb{R}_+})$  equipped with a standard Brownian motion B and a random Poisson measure N.

► (H1) - Regularity. We assume that there exists constants  $L_a$ ,  $L_b$  and a measurable function  $L_c: [0,T] \times \mathbb{R} \to \mathbb{R}_+$ 

$$\begin{aligned} |b(x) - b(y)| &\leq L_b \ (|x - y|), & x, y \in \mathbb{R}, \\ |c(t, x, z) - c(t, y, z)| &\leq L_c(t, z) \ |x - y|, & x, y \in \mathbb{R}, t \in [0, T], z \in \mathbb{R}. \end{aligned}$$

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• (H2) - Integrability. We assume that the function  $\psi_p$  defined by

$$\psi_p(t) = \left(\int_{-\infty}^{\infty} |L_c(t,z) \vee |c(t,0,z)||^2 v_t(dz)\right)^{p/2} + \int_{-\infty}^{\infty} |L_c(t,z) \vee |c(t,0,z)||^p v_t(dz)$$

belongs to  $L^{1+\zeta}([0,T])$  for some  $\zeta \in (0,1]$ .

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- ▶ (H3) A.R. Approximation. We assume that there exists  $\varepsilon^* \in (0,1]$  such that
  - (Moments) for all  $x \in \mathbb{R}$  and  $t \in [0, T]$ ,  $c(t, x, \cdot)$  is not identically zero on  $B(\varepsilon^*)$  and

$$\int_0^T \int_{B(e^*)} |c(t,x,z)|^{p+2} v_t(dz) dt + \int_0^T \left( \int_{B(e^*)} |c(t,x,z)|^2 v_t(dz) \right)^{\frac{p}{2}+1} dt < \infty.$$

(Coupling) for all x ∈ ℝ and t ∈ [0, T], the image measure of 1 {z∈B(ε\*)} v<sub>t</sub>(dz) by z → c(t,x,z) has a density with respect to the Lebesgue measure on ℝ and satisfies ∫<sub>B(ε\*)</sub> v<sub>t</sub>(dz) = ∞.

## 5. Strong convergence of $\overline{X}^{\varepsilon}$ in $L^p(\Omega)$ : main theorem

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#### Théorème 1 – (Bossy, Maurer 2023)

We assume that (H1), (H2) and (H3) hold.

(*i.*) For any  $\varepsilon \in (0, \varepsilon^*]$ , there exists a sequence  $(\widehat{X}_{t_l}^{\varepsilon})_{i \in \{0,...,n\}}$  of random variables on  $(\Omega, \mathcal{F})$ , such that for any  $i \in \{0, ..., n\}$ ,  $\overline{X}_{t_l}^{\varepsilon}$  is  $\mathcal{F}_{t_l}$ -measurable, and verifies  $\operatorname{Law}(\widehat{X}_{t_l}^{\varepsilon}) = \operatorname{Law}(\overline{X}_{t_l}^{\varepsilon})$ . Moreover, there exists  $\overline{m}(p, T) > 0$  such that

$$\sup_{\varepsilon \in [0,\varepsilon^*)} \mathbb{E} \Big[ \sup_{i \in \{0,\dots,n\}} |\widehat{X}_{t_i}^{\varepsilon}|^p \Big] \le \overline{m}(p,T).$$

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(*ii*.) The following inequality stands true for any  $\varepsilon \in (0, \varepsilon^*]$ :

$$\left\|\sup_{i\in\{0,\dots,n\}}\left|X_{t_i}-\widehat{X}_{t_i}^{\varepsilon}\right|\right\|_{L^p(\Omega)} \preccurlyeq n^{-\left\{\frac{2\zeta}{p(1+\zeta)}\right\}}+\delta_p^n(\varepsilon),$$

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where 
$$\delta_p^n(\varepsilon) = \left(\sum_{k=1}^n \mathbb{E}\left[\left(\frac{\int_{l_{k-1}}^{l_k} \int_{B(\varepsilon)} |c(s, \widehat{X}_{l_{k-1}}^{\varepsilon}, z)|^{p+2} \mathbf{v}_s(dz) ds}{\int_{l_{k-1}}^{l_k} \int_{B(\varepsilon)} |c(s, \widehat{X}_{l_{k-1}}^{\varepsilon}, z)|^2 \mathbf{v}_s(dz) ds}\right)^{\frac{1}{p}}\right]^2\right)^{\frac{1}{2}}$$

satisfies  $\lim_{\epsilon\to 0} \delta_p^n(\epsilon)=0$  when the following sufficient condition holds:

 $\lim_{|z|\to 0} \sup_{t\in[0,T]} |L_c(t,z)| = 0.$ 

#### Corollaire

We assume in addition to (H1), (H2) and (H3) that there exists a constant  $C_T$  satisfying

$$\forall (s,x,z) \in [0,T] \times \mathbb{R} \times B(\boldsymbol{\varepsilon}^*), \quad |c(t,x,z)| \le C_T |z| (1+|x|).$$

$$\tag{1}$$

Then, for any  $\varepsilon \in (0, \varepsilon^*]$ , the *L*<sup>*p*</sup>-strong error of the  $((\Omega, \mathcal{F})$ -representation of the)  $\varepsilon$ -EM scheme  $\widehat{X}^{\varepsilon}$  satisfies:

$$\Big|\sup_{i\in\{0,\dots,n\}}|X_{t_i}-\widehat{X}_{t_i}^{\varepsilon}|\Big\|_{L^p(\Omega)}\preccurlyeq n^{-\left\{\frac{2\zeta}{p(1+\zeta)}\right\}}+\varepsilon\sqrt{n}.$$

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▶ Moreover, suppose we have  $\psi_p \in L^2([0,T])$  (i.e.  $\zeta = 1$ ). With  $\varepsilon$  taken such that

$$\varepsilon \leq n^{-\left(\frac{1}{2}+\frac{1}{p}\right)} \wedge \varepsilon^*,$$

we obtain the following convergence rate for the  $L^p$ -strong error:

$$\left\|\sup_{i\in\{0,\dots,n\}}\left|X_{t_i}-\widehat{X}_{t_i}^{\varepsilon}\right|\right\|_{L^p(\Omega)} \preccurlyeq n^{-\frac{1}{p}}$$

#### The continuous Euler-Peano scheme as a pivot term

We use as a pivot term the SDE with frozen coefficients (or Euler-Peano scheme) X defined by

$$\widetilde{X}_t = \int_0^t b(\widetilde{X}_{\eta(s)}) ds + \int_0^t \int_{-\infty}^{+\infty} c(s, \widetilde{X}_{\eta(s^-)}, z) \widetilde{N}(ds, dz),$$

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• We first prove a rate of convergence for  $\widetilde{X}$  and then compare  $\widetilde{X}$  with our scheme:

#### Proposition 2 - L<sup>p</sup>-convergence of the Euler-Peano scheme

Assume (H1) and (H2). Then for all  $n \in \mathbb{N}^*$ ,

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The proof of Proposition 2 relies on a Gronwall argument as it is usually the case for a standard strong convergence proof, but has some specificity due to the non-continuous paths of the process and the time-inhomogeneity of the jumps.

#### Kunita inequality

To prove Proposition 2, we need a tool to estimate the L<sup>p</sup>-moments of a stochastic Poisson integral. Since these types of integrals are martingales, one may want to use the Burkhölder-Davis-Gundy (BDG) inequality :

$$\mathbb{E}\left[\sup_{0\leq s\leq t}|M_s|^p\right]\leq C_p^{\textit{BDG}}\,\mathbb{E}[[M,M]_t^{p/2}],$$

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- However, in the case of discontinuous process, the quadratic variation [M, M] is not equal to the predictable quadratic variation (M, M), and we only have tools to estimate the latter.
- For this reason, we use another inequality that is more specific to Poisson integrals:

#### Lemma (Kunita inequality)

Let *F* be a predictable stochastic process and  $I_t = \int_0^t \int_{-\infty}^{\infty} F(s, z) \widetilde{N}(ds, dz)$ . Then for all  $p \ge 2$  there exists a constant *C* depending only on *p* and *T* such that

$$\mathbb{E}\left[\sup_{s\in[0,t]}|I_s|^p\right] \le C\int_0^t \mathbb{E}\left[\left(\int_{-\infty}^{\infty}|F(s,z)|^2\nu_s(dz)\right)^{p/2}\right]ds + C\int_0^t \mathbb{E}\left[\int_{-\infty}^{\infty}|F(s,z)|^p\nu_s(dz)\right]ds$$

#### Proof of proposition 2 - The Brownian case

- Let's first recall the scheme of the proof in the standard (Brownian noise) case.
- Setting ℰ(t) = sup<sub>s∈[0,t]</sub> ||X<sub>s</sub> − X̃<sub>s</sub>||<sub>L<sup>p</sup>(Ω)</sub>, we may use Minkoswki integral inequality and BDG inequality to obtain the upper-bound

$$\mathscr{E}(t) \preccurlyeq \int_0^t \|b(X_s) - b(\widetilde{X}_{\eta(s)})\|_{L^p(\Omega)} ds + \left(\int_0^t \|\sigma(X_s) - \sigma(\widetilde{X}_{\eta(s)})\|_{L^p(\Omega)}^2 ds\right)^{\frac{1}{2}}.$$

Then we use the Lipschitz property of b and σ to get

$$\mathscr{C}(t) \preccurlyeq \int_0^t \|X_s - \widetilde{X}_{\eta(s)}\|_{L^p(\Omega)} ds + \left(\int_0^t \|X_s - \widetilde{X}_{\eta(s)}\|_{L^p(\Omega)}^2 ds\right)^{\frac{1}{2}}.$$

We may use the pivot X<sub>η(s)</sub> to separate the two terms of the right-hand side into a local error term and a Gronwall term, as follows:

$$\mathscr{E}(t) \preccurlyeq \int_0^t (\|X_s - X_{\eta(s)}\|_{L^p(\Omega)} + \mathscr{E}(s))ds + \left(\int_0^t (\|X_s - X_{\eta(s)}\|_{L^p(\Omega)}^2 + \mathscr{E}(s)^2)ds\right)^{\frac{1}{2}}.$$

Finally we may bound the local error terms using the same inequalities and s − η(s) ≤ <sup>1</sup>/<sub>n</sub>, allowing to apply a Gronwall-type lemma. This gives a rate of convergence of n<sup>-<sup>1</sup>/<sub>2</sub></sup>, where the power 1/2 comes from the BDG inequality.

#### Proof of proposition 2 - The Poisson case

- We now move to the (sketch of the) proof in our case.
- Setting  $\mathscr{C}(t) = \sup_{s \in [0,t]} ||X_s \widetilde{X}_s||_{L^p(\Omega)}$ , we may use Minkoswki integral inequality and Kunita inequality to obtain the upper-bound

$$\begin{aligned} \mathscr{C}(t) \preccurlyeq &\int_{0}^{t} \|b(X_{s}) - b(\widetilde{X}_{\eta(s)})\|_{L^{p}(\Omega)} ds + \left(\int_{0}^{t} \mathbb{E}\left[\left(\int_{-\infty}^{\infty} (c(s,X_{s},z) - c(s,\widetilde{X}_{\eta(s)},z))^{2} v_{s}(dz)\right)^{\frac{p}{2}}\right] ds\right)^{\frac{1}{p}} \\ &+ \left(\int_{0}^{t} \int_{-\infty}^{\infty} \mathbb{E}\left[|c(s,X_{s},z) - c(s,\widetilde{X}_{\eta(s)},z)|^{p}\right] v_{s}(dz) ds\right)^{\frac{1}{p}}.\end{aligned}$$

Then we use the Lipschitz property of b and c to get

$$\mathscr{C}(t) \preccurlyeq \int_0^t \|X_s - \widetilde{X}_{\eta(s)}\|_{L^p(\Omega)} ds + \left(\int_0^t \psi_p(s) \left\|X_s - \widetilde{X}_{\eta(s)}\right\|_{L^p(\Omega)}^p ds\right)^{\frac{1}{p}}$$

where we recall that:

$$\psi_p(t) = \left(\int_{-\infty}^{\infty} |L_c(t,z) \vee |c(t,0,z)||^2 v_t(dz)\right)^{p/2} + \int_{-\infty}^{\infty} |L_c(t,z) \vee |c(t,0,z)||^p v_t(dz)$$

Sketch of the proof of Theorem 1. To prove Theorem 1, we need to do three things.

- Construct a version X<sup>ε</sup> of the scheme that lives on the same probability space as X and satisfy an optimal coupling property with respect to the *W<sub>p</sub>*-distance: we will detail this construction on the next slide.
- 2. Find a **uniform bound** in  $\varepsilon$  for  $\hat{X}^{\varepsilon}$ : this is done by a standard Gronwall argument using some discrete martingale properties of the construction of  $\hat{X}^{\varepsilon}$  and the discrete BDG inequality.
- Derive an upper-bound for the L<sup>p</sup>-norm of X̃ − X<sup>e</sup>: this is also done by a Gronwall argument using the optimal coupling property and Proposition 1 to get an upper bound for the small jump approximation part, leading to the contribution δ<sup>n</sup><sub>n</sub>(ε) in the result of Theorem 1.

#### More details on the construction of $\widehat{X}^{\varepsilon}$

For fixed *x*, we denote  $Y_i(x) = \int_{t_{i-1}}^{t_i} \int_{B(\varepsilon)} c(s, x, z) \widetilde{N}(ds, dz)$ . Our aim consists in constructing a map  $T_i : \mathbb{R} \times \mathbb{R} \to \mathbb{R}$  that optimally transports (for the  $\mathcal{W}_p$ -distance) Law $(Y_i(x))$  to the centred normal distribution  $\mathcal{N}_i(x)$  of variance  $\int_{t_{i-1}}^{t_i} \int_{B(\varepsilon)} c(s, x, z)^2 v_s(dz) ds$ . In addition, this map must be measurable with respect to the parameter *x*.

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- ▶ We set  $\widehat{X}_{t_0}^{\varepsilon} = X_0$ . For  $i \in \{1, ..., n\}$ , let  $\mathbb{Q}_i = \text{Law}(\overline{X}_{t_{i-1}}^{\varepsilon})$ . Applying Theorem 1.1 from Fontbona-Guérin-Meleard (see [FGM10]), there exists an application  $T_i : \mathbb{R} \times \mathbb{R} \to \mathbb{R}$  which is  $(\mathscr{B}(\mathbb{R}) \otimes \mathscr{B}(\mathbb{R}), \mathscr{B}(\mathbb{R}))$ -measurable such that for  $\mathbb{Q}_i$ -almost every  $x \in \mathbb{R}$ , one has

$$\mathbb{E}[|Y_i(x) - T_i(x, Y_i(x))|^p] = \mathcal{W}_p(\mathsf{Law}(Y_i(x)), \mathcal{N}_i(x))^p.$$

• Then, given  $\widehat{X}_{t_{i-1}}^{\varepsilon}$ , we can set

$$\widehat{X}_{t_i}^{\varepsilon} = \widehat{X}_{t_{i-1}}^{\varepsilon} + b(\widehat{X}_{t_{i-1}}^{\varepsilon})(t_i - t_{i-1}) + T_i\left(\widehat{X}_{t_{i-1}}^{\varepsilon}, Y_i(\widehat{X}_{t_{i-1}}^{\varepsilon})\right) + \int_{t_{i-1}}^{t_i} \int_{\mathbb{R}\setminus B(\varepsilon)} c(s, \widehat{X}_{t_{i-1}}^{\varepsilon}, z)\widetilde{N}(ds, dz).$$

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• Then, given  $\widehat{X}_{t_{i-1}}^{\varepsilon}$ , we can set

$$\widehat{X}_{t_i}^{\varepsilon} = \widehat{X}_{t_{i-1}}^{\varepsilon} + b(\widehat{X}_{t_{i-1}}^{\varepsilon})(t_i - t_{i-1}) + T_i\left(\widehat{X}_{t_{i-1}}^{\varepsilon}, Y_i(\widehat{X}_{t_{i-1}}^{\varepsilon})\right) + \int_{t_{i-1}}^{t_i} \int_{\mathbb{R}\setminus B(\varepsilon)} c(s, \widehat{X}_{t_{i-1}}^{\varepsilon}, z)\widetilde{N}(ds, dz).$$

By construction, X<sup>ε</sup> and X<sup>ε</sup> have the same law, and one can check that (X<sup>ε</sup><sub>i1</sub>, 0 ≤ i ≤ n) is an adapted sequence to the filtration (𝔅<sub>i1</sub>, 0 ≤ i ≤ n), and that (T<sub>i</sub>(X<sup>ε</sup><sub>i1-1</sub>, Y<sub>i</sub>(X<sup>ε</sup><sub>i1-1</sub>)), 1 ≤ i ≤ n) is a sequence of (discrete) martingale increments relatively to this filtration.

There are two limitations to a numerical evaluation of the strong convergence rate in our case that we want to point out:

#### 1. The lack of exact trajectory solution.

Processing the computation of the strong norm of the error always poses the problem of simulating a reference solution trajectory, that is not available in the jump case. We chosed to compute an **approximate reference solution**, by pushing the approximation parameters to a limit value which serves as a bound for the experiments with coarse parameters. This forces us to restrict the numerical test in the increment case  $c(s, x, z) = \sigma(x)f(s, z)$ .

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#### 2. The sampling of the two-parameters increments.

For the  $\varepsilon$ -EM algorithm, we have two control-parameters, the time step 1/n and the small jumps cut  $\varepsilon$ . Once the choice of  $\varepsilon$  is **fixed**, the increments of the process  $\int_{0}^{1} \int_{\mathbb{R}/B(\varepsilon)} z\widetilde{N}(ds, dz)$  can then be simulated on a very fine time grid, and next aggregated together to produce increments on a coarser time grid.

#### Rate of convergence in terms of the norm exponent p

We investigate the behaviour of the  $L^p$ -strong error with respect to the variations of  $p \ge 2$ . For that purpose, we consider the following example:

$$X_t = \int_0^t \cos(X_s) ds + \int_0^t \int_{-\infty}^\infty \sin(X_{s^-}) \ z \ \widetilde{N}(ds, dz), \quad \mathbf{v}_s(dz) ds = \mathbb{1}_{\{|z| \le 10\}} \frac{dz}{|z|^{3/2}} \ ds.$$



#### Rate of convergence with low time-integrability

When  $\psi_p$  is not more than  $L^{1+\zeta}$  with  $\zeta \in (0,1)$ , we may only recover the rate  $n^{-\frac{2\zeta}{p(1+\zeta)}}$ . We investigated if this loss could be observed numerically with the following equation:

$$X_{t} = \int_{0}^{t} \sin(X_{s}) ds + \int_{0}^{t} \int_{-\infty}^{\infty} \cos(X_{s-}) \ z \widetilde{N}(ds, dz), \quad v_{s}(dz) ds = \mathbb{1}_{\{|z| \le 10\}} \frac{dz}{|z|^{3/2}} \ s^{\beta} ds, \quad \beta \in (-1, 0]$$



Figure 1: Behaviour of  $\left\|\sup_{i \in \{0,...,n\}} \left| \overline{X}_{t_i}^{\varepsilon_{\min},n} - \overline{X}_{t_i}^{\varepsilon_{\min},n} - \overline{X}_{t_i}^{\varepsilon_{\min},n} \right| \right\|_{L^p(\Omega)}$  with n, for various  $L^p$ -norms (lines with makers), and the corresponding theoretical (dash lines) rates.

### 9. Conclusion

- We developed a numerical scheme to approximate a class of time-inhomogeneous jump SDEs based on the Asmussen-Rosinski technique, and derived a rate of convergence for the L<sup>p</sup>-strong error by optimal transport technique, using bounds for the CLT convergence in W<sub>p</sub>-distance.
- In the setting we are and assuming more regularity on the coefficients, we believe that it is possible to also derive a weak error rate for this scheme. This work is still in progress, but we have good confidence to obtain the following result:

#### Theorem (Work in progress)

Assume (H1) and "good enough" space and time regularity of the coefficients. Assume that the v<sub>t</sub> are dominated by a Lévy measure  $\mu$ . Let  $\beta = \inf\{\alpha > 0, \int_{|z|<1} |z|^{\alpha} \mu(dz) < \infty\}$  the Blumenthal-Getoor index of the indivisible distribution characterised by  $\mu$ . Let  $\phi \in \mathfrak{C}^4(\mathbb{R})$  be such that for every  $k \in \{1, \ldots, 4\}$ ,

$$\left. \frac{\partial^k \varphi}{\partial x^k}(x) \right| \le C(1+|x|^q) \tag{2}$$

for some  $q \leq \frac{p}{2}$ . We have the following weak error upper-bound:

$$|\mathscr{C}[\varphi(X_T^{\varepsilon})] - \mathscr{C}[\varphi(\overline{X}_T^{\varepsilon})]| \preccurlyeq n^{-1} + \varepsilon^{3-\beta^+}.$$

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